

Article

# An alternative approach to approximating moments of least squares estimators

Gareth Liu-Evans

Management School, University of Liverpool, Chatham Street, Liverpool, United Kingdom

Version April 19, 2015 submitted to *Econometrics*. Typeset by  $\LaTeX$  using class file *mdpi.cls*

---

**Abstract:** Results are presented for approximating the moments of least squares estimators, particularly those of the OLS estimator, and the methodology is illustrated using a simple dynamic model. An  $O(T^{-1})$  approximation is presented for the bias in estimation of a general ARX( $p$ ) model, and this is specialised to known cases for the AR(1) with and without constant under an assumption of non normal model disturbances.

**Keywords:** asymptotic approximation; bias; least squares ; time series ; autoregression ; simulteneity

---

## 1. Introduction

We present results to facilitate the asymptotic approximation of the moments of least squares coefficient estimators under similar assumptions to [1], but focussing on the OLS estimator. The expansion method in [1] is valid for consistent  $k$ -class estimation of equations from static simultaneous equation systems, but the methodology does not apply to OLS or dynamic models. We instead use the validity framework in [2], and results are presented here to aid the direct use of the approach with autoregressive models as an illustration.

This is done using matrix differential calculus results from [3], [4], and [5,6]. The approach builds on [7] and [8], who consider the moments of the  $k$ -th order serial correlation coefficient in AR(1) models. It is shown how the approach is applicable in principle to models where both dynamics and simulteneity are present, as an alternative to methodology based on [9].

There are large- $T$  asymptotic approximations of the OLS estimator moments using [9]-type methodologies in [10,11] and [12] in the context of stationary autoregressive models, and in [13] for equations from static simultaneous equation models. Building on [14], [15] present a general method

22 for estimators of time-series models. See also [16,17] and [18].

23

## 24 2. The expansion method

25 Given a model

$$y = Z\alpha + u \quad (1)$$

26 with  $E[Z'u] = 0$  (this assumption is dropped later), and with  $Z$  being a  $T \times N$  matrix with rank  $N$   
 27 almost surely, the true coefficient  $\alpha$  and its OLS estimator  $\hat{\alpha}$  can be expressed in the same functional  
 28 form so long as the expected values in (2) exist:

$$\hat{\alpha} = (Z'Z)^{-1}Z'y \quad \text{and} \quad \alpha = (E[Z'Z])^{-1}E[Z'y]. \quad (2)$$

29 Equation (3) is obtained by premultiplying both sides of (1) by  $Z'$ . By defining matrices  $\hat{R} = [Z'Z : \hat{\zeta}]$   
 30 and  $R = [E[Z'Z] : \zeta]$ , where  $\hat{\zeta} = Z'y$  and  $\zeta = E[Z'y]$ , the estimated and true coefficients can be  
 31 expressed as

$$\hat{\alpha}_i = f_i(\hat{\delta}) \quad \text{and} \quad \alpha_i = f_i(\delta), \quad (3)$$

32 respectively, for  $i = 1, \dots, N$ , where  $\hat{\delta} = \text{vec}(\hat{R})$  and  $\delta = \text{vec}(R)$ . This allows a Taylor series  
 33 expansion of the following form:

$$\hat{\alpha}_i - \alpha_i = (\hat{\delta} - \delta)' f_i'(\delta) + \frac{1}{2}(\hat{\delta} - \delta)' H_i|_{\delta} (\hat{\delta} - \delta) + \dots, \quad (4)$$

34 where  $H_i|_{\delta}$  is the Hessian matrix of  $f$  evaluated at  $\delta = \text{vec}(R)$ .

35 Using the extended mean value theorem we may write

$$\begin{aligned} f_i(\hat{\delta}) &= f_i(\delta) + (\hat{\delta} - \delta)' f_i'(\delta) + \frac{1}{2}(\hat{\delta} - \delta)' H_i|_{\delta} (\hat{\delta} - \delta) \\ &\quad + \frac{1}{3!} \sum_{j=1}^r (\hat{\delta}_j - \delta_j) (\hat{\delta} - \delta)' f_{ij}^{(3)}|_{\delta^*} (\hat{\delta} - \delta) \end{aligned} \quad (5)$$

36 for some  $\delta^*$ , where  $r$  denotes the row dimension of  $\delta$ , and where  $f_{ij}^{(3)}$  is an  $r \times r$  matrix of derivatives  
 37 defined as  $f_{ij}^{(3)} = \frac{\partial^3 H_i}{\partial \delta_j^3}$ . Here it is assumed that  $f_i$  is differentiable up to third order and that the derivatives  
 38 are uniformly bounded in a neighbourhood of  $\delta$  as  $T \rightarrow \infty$ , and with third-order derivatives that are  
 39 continuous. This makes the fourth term  $O_p(T^{-\frac{3}{2}})$ . If the components of  $\hat{\delta}$  are assumed to have finite  
 40 moments up to third order we therefore have

$$E[\hat{\alpha}_i - \alpha_i] = \frac{1}{2} \text{tr}(H_i|_{\delta} \text{Var}(\hat{\delta})) + o(T^{-1}), \quad (6)$$

41 where  $Var(\hat{\delta})$  is the covariance matrix for  $\hat{\delta}$ . This final step is clear from [19] section 2.4 and [20],  
 42 see also [2], Appendix A. Rearranging this slightly we have

43 **Theorem 1.** *Under the assumptions made,*

$$E[\hat{\alpha}_i - \alpha_i] = \frac{1}{2} (tr(H_i|_{\delta}J) + \delta' H_i|_{\delta} \delta) + o(T^{-1}),$$

44 where  $J = E[\hat{\delta}\hat{\delta}']$ .

45 Writing  $\Gamma_1 = [0_{N \times N^2} : I_N]$ ,  $\Gamma_2 = [I_{N^2} : 0_{N^2 \times N}]$ ,  $V_1 = vec(I_N)$ ,  $V_2 = (K_{NN} \otimes I_N)(I_N \otimes vec(Z'Z))$   
 46 and  $V_3 = (I_N \otimes K_{NN})(vec(Z'Z) \otimes I_N)$ , where  $K_{nm}$  is an  $nm \times nm$  commutation matrix and  $I_N$  is  
 47 an  $N \times N$  identity matrix, Theorem 2 below provides the Hessian matrix  $H_i$  in terms of  $\hat{\zeta}$  and  $Z'Z$ ,  
 48 using the Second Identification Theorem in [6]. To evaluate the Hessian at  $\delta$ , we replace  $\hat{\zeta}$  and  $Z'Z$   
 49 with their expected values, and this is done later for autoregressive models. In the following we write  
 50  $A \otimes A \otimes \dots \otimes A = A^{\otimes m}$ , where  $\otimes$  is the Kronecker product and  $A$  appears  $m$  times. Since  $(A \otimes A \otimes$   
 51  $\dots \otimes A)^{-1} = A^{-1} \otimes A^{-1} \otimes \dots \otimes A^{-1}$  when  $A$  is invertible, we denote this by  $A^{\otimes(-m)}$ .

52 **Theorem 2.** *The Hessian matrix is*

$$H_i = \frac{1}{2} (MB_i + B_i'M'),$$

53 where

$$M = \left( \begin{aligned} & ([\Gamma_2' \otimes (\hat{\zeta}' \otimes I_N)](Z'Z)^{\otimes(-4)}[(I_N \otimes V_2) + (V_3 \otimes I_N)]\Gamma_2 \\ & - (\Gamma_1' \otimes I_N)(Z'Z)^{\otimes(-2)}\Gamma_2 - [((Z'Z)^{\otimes(-2)}\Gamma_2)' \otimes I_N](I_N \otimes V_1)\Gamma_1 \end{aligned} \right)',$$

54 and where  $B_i$  is defined as

$$\begin{aligned} (B_i)_{n,m} &= 1 \text{ for } n = qN + i, \quad m = q + 1, \quad q = 0, 1, \dots, N^2 + N - 1 \\ &= 0 \text{ otherwise.} \end{aligned}$$

55 *Proof.* See Appendix A.

56 It may be possible to extend the methodology here to cases where  $E[Z'u] \neq 0$ , i.e. where endogeneity  
 57 is present. We have the following by premultiplying  $y = Z\alpha + u$  by  $Z'$  and taking the expected value as  
 58 above:

$$\begin{aligned} \alpha &= (E[Z'Z])^{-1}E[Z'y] - (E[Z'Z])^{-1}E[Z'u] \\ \Rightarrow \alpha_i &= f_i(\delta) - e_i'E_1^{-1}E_2, \end{aligned}$$

59 where  $E_1 = E[Z'Z]$  and  $E_2 = E[Z'u]$ , and where  $e_i$  is an  $N \times 1$  unit vector with unity in position  $i$ .  
60 Since it is still true that

$$E[\hat{\alpha}_i] = f_i(\delta) + \frac{1}{2} (tr(H_i|_\delta J) + \delta' H_i|_\delta \delta) + o(T^{-1}),$$

61 the bias in OLS estimation is

$$E[\hat{\alpha}_i - \alpha_i] = \frac{1}{2} (tr(H_i|_\delta J) + \delta' H_i|_\delta \delta) + e_i' E_1^{-1} E_2 + o(T^{-1}),$$

62 and the approximation is valid to order  $O(T^{-1})$  under the same conditions as before.

63 The following sections apply Theorem 1 to the ARX( $p$ ) with non-normal disturbances. As in [15]  
64 and [12] the third and fourth moments of the model errors are expressed in terms of skewness and excess  
65 kurtosis parameters, so that the effects of departures from normality can be seen more easily.

### 67 3. ARX( $p$ ) illustration

68 Consider an autoregressive model with  $p$  lags and  $k$  added exogenous variables:

$$y = \lambda_1 y_{-1} + \dots + \lambda_p y_{-p} + X\beta + u,$$

69 where  $u = \Gamma_3 v$  with  $\Gamma_3 = [0_{T \times p} : I_T]$ , and where  $v$  is a  $(T+p) \times 1$  random vector with the following  
70 moment properties.

71 **Assumption 1.** *The  $i$  – th elements of  $v$  have finite moments up to 6th order with:*

$$E[v_i] = 0, \quad E[v_i^2] = \sigma^2 \quad E[v_i^3] = \sigma^3 \gamma_1 \quad E[v_i^4] = \sigma^4 (\gamma_2 + 3),$$

72 where  $\gamma_1$  and  $\gamma_2$  are Pearson's measures of skewness and excess kurtosis.

73 It is also assumed that the process is stationary in the sense that all roots of  $1 - \lambda_1 r - \lambda_2 r^2 - \dots - \lambda_p r^p = 0$   
74 lie outside the unit circle. This assumption, combined with Assumption 1, makes the process covariance  
75 stationary. The assumption of finite moments to 6th order for  $v$  ensures that  $\hat{\delta}$  has finite moments up to  
76 3rd order, which is a condition for Theorem 1.

77

78 From this we can write the following for periods  $1 - p$  through to  $T - 1$ , building on the approach in  
79 [11]<sup>1</sup>:

$$\Lambda Y_{-1} = \bar{Y}_{-1}^* + [I_{T+p-1} : 0_{(T+p-1) \times 1}] \Omega v,$$

---

<sup>1</sup> See also [21].

80 where  $Y_{-1} = (y_{1-p}, \dots, y_{T-1})'$ ,  $\bar{Y}_{-1}^* = (\bar{y}_{1-p}, \dots, \bar{y}_0, x_1'\beta, \dots, x_T'\beta)'$  and  $\bar{Y}_{-1} = \Lambda^{-1}\bar{Y}_{-1}^*$ . The matrices  
81  $\Lambda$  and  $\Omega$  are defined, respectively, as

$$\Lambda = \begin{pmatrix} 1 & 0 & \cdot & \cdot & \cdot & \cdot & 0 \\ 0 & 1 & 0 & \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot & \cdot & \cdot & \cdot \\ 0 & \cdot & 0 & \cdot & \cdot & \cdot & \cdot \\ -\lambda_p & \cdot & \cdot & -\lambda_1 & \cdot & 0 & \cdot \\ 0 & \cdot & \cdot & \cdot & \cdot & 1 & 0 \\ 0 & 0 & -\lambda_p & \cdot & \cdot & -\lambda_1 & 1 \end{pmatrix} \text{ and } \Omega = \begin{pmatrix} \omega & 0 & \cdot & \cdot & \cdot & \cdot & 0 \\ 0 & \cdot & 0 & \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \omega & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot & 1 & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot & \cdot & \cdot & 0 \\ 0 & \cdot & \cdot & \cdot & \cdot & 0 & 1 \end{pmatrix},$$

82 where  $\Lambda$  has  $\lambda_{-p}$  as the  $(p+1)th$  element of the first column, and where the  $\omega$  term, which is the  
83 standard deviation of  $y_t$ , appears  $p$  times in  $\Omega$ . This can alternatively be written as

$$Y_{-1} = \bar{Y}_{-1} + Gv,$$

84 where  $G = \Lambda^{-1}[I_{T+p-1} : 0_{(T+p-1) \times 1}]$ , and we can note that  $y_{-i} = M_i Y_{-i}$  and  $\bar{y}_{-i} = M_i \bar{Y}_{-i}$  for  
85  $i = 1, \dots, p$ , where  $M_i = [0_{T \times (p-i)} : I_T : 0_{T \times (i-1)}]$ . This implies

$$y_{-i} = \bar{y}_{-i} + G_i v,$$

86 where  $G_i = M_i G$ . When  $p = 1$ ,  $M_i$  is the identity matrix and  $G$  is the same  $G$  that appears in [11].

87 Following [11] we write  $Z = \bar{Z} + \tilde{Z}$ , where  $\bar{Z} = (\bar{y}_{-1}, \dots, \bar{y}_{-p}, X)$  and  $\tilde{Z} = (G_1 v, \dots, G_p v, 0_{T \times k})$ .

88 We have  $\tilde{Z} = (G_1 v, \dots, G_p v) \Gamma_4$  where  $\Gamma_4 = [I_p : 0_{p \times k}]$ , which gives

$$Z = \bar{Z} + \sum_{i=1}^p G_i v e_i' \Gamma_4,$$

89 where  $e_i$  here is a  $p \times 1$  vector with 1 in position  $i$  and 0 elsewhere. When  $p = 1$ , the result reduces  
90 to  $Z = \bar{Z} + G v e_1'$ , which appears in [11].

91 We may write  $\hat{\delta}$  as a linear combination of  $v$  and quadratic forms in  $v$ , so that the expected value of  
92  $\hat{\delta} \hat{\delta}'$  can be calculated using existing results on the expected values of products of quadratic forms.<sup>2</sup> We  
93 note that  $\hat{\delta} = \Gamma_2' \text{vec}(Z'Z) + \Gamma_1' \text{vec}(\hat{\zeta})$ , where  $\Gamma_1$  and  $\Gamma_2$  are the same as in the preceding section with  
94  $N = p + k$ , and then express  $\text{vec}(Z'Z)$  and  $\text{vec}(\hat{\zeta})$  in terms of  $v$ .

95

96 We may write  $y$  in terms of  $v$ :

---

<sup>2</sup> See e.g. the Appendix in [22].

$$\begin{aligned}
y &= \sum_{i=1}^p \lambda_i (\bar{y}_{-i} + G_i v) + X\beta + u \\
&= \left( \sum_{i=1}^p \lambda_i \bar{y}_{-i} \right) + X\beta + \left\{ \left( \sum_{i=1}^p \lambda_i G_i \right) + \Gamma_3 \right\} v.
\end{aligned}$$

97 Using these decompositions of  $y$  and  $Z$ , it is straightforward to express  $\text{vec}(\hat{\zeta})$  and  $\text{vec}(Z'Z)$  in the  
98 desired form, and this is done in Lemma 1 below.

99 **Lemma 1.** *In the ARX( $p$ ) model the terms  $\text{vec}(\hat{\zeta})$  and  $\text{vec}(Z'Z)$  are*

$$\begin{aligned}
\text{vec}(\hat{\zeta}) &= P_1 + P_2 v + \sum_{i=1}^p P_{3i} v' P_{4i} v \\
\text{vec}(Z'Z) &= A_1 + A_2 v + \sum_{i,j=1}^p A_{3ij} v' A_{4ij} v,
\end{aligned}$$

100 *where*

$$\begin{aligned}
P_1 &= \text{vec}[\bar{Z}' \{ (\bar{Z} \sum_{i=1}^p \lambda_i \bar{y}_{-i}) + X\beta \}] \\
P_2 &= \bar{Z}' \{ (\sum_{i=1}^p \lambda_i G_i) + \Gamma_3 \} + \sum_{i=1}^p \{ [ (\sum_{j=1}^p \lambda_j \bar{y}_{-j}) + X\beta ]' G_i \} \otimes \Gamma_4' e_i \\
P_{3i} &= \text{vec}(\Gamma_4' e_i), \quad P_{4i} = G_i' \{ (\sum_{j=1}^p \lambda_j G_j) + \Gamma_3 \}, \quad A_1 = \text{vec}(\bar{Z}' \bar{Z}) \\
A_2 &= \sum_{i=1}^p \{ (\Gamma_4' e_i \otimes \bar{Z}' G_i) + (\bar{Z}' G_i \otimes \Gamma_4' e_i) \}, \quad A_{3ij} = \Gamma_4' e_j \otimes \Gamma_4' e_i, \quad A_{4ij} = G_i' G_j
\end{aligned}$$

101 *Proof. See Appendix B.1.*

102 Using Lemma 1, it is possible to calculate the expected value  $J$ , and this is provided in Lemma 2.

103 **Lemma 2.** *The expected value  $J$  is*

$$\begin{aligned}
J &= Q_1 Q_1' + \sigma^2 \sum_{i,j=1}^p \text{tr}(Q'_{4ij}) Q_1 Q'_{3ij} + \sigma^2 \sum_{i=1}^p \text{tr}(Q'_{6i}) Q_1 Q'_{5i} + \sigma^2 Q_2 Q_2' \\
&+ \sigma^2 \sum_{i,j=1}^p \text{tr}(Q_{4ij}) Q_{3ij} Q_1' + \sigma^4 \sum_{i,j,k,l=1}^p Q_{3ij} \text{tr}[Q'_{4lm} \{ \text{tr}(Q_{4ij}) I_{T+1} + Q_{4ij} + Q'_{4ij} \}] Q'_{3lm} \\
&+ \sigma^4 \sum_{i,j,l=1}^p Q_{3ij} \text{tr}[Q'_{6l} \{ \text{tr}(Q_{4ij}) I_{T+1} + Q_{ij4} + Q'_{4ij} \}] Q'_{5l} \\
&+ \sigma^2 \sum_{i=1}^p \text{tr}(Q_{i6}) Q_{5i} Q_1' + \sigma^4 \sum_{i,j,l=1}^p Q_{5i} \text{tr}[Q'_{4jl} \{ \text{tr}(Q_{6i} I_{T+1} + Q_{6i} + Q'_{6i}) \}] Q'_{3jl} \\
&+ \sigma^4 \sum_{i,j=1}^p Q_{5i} \text{tr}[Q'_{6j} \{ \text{tr}(Q_{6j}) I_{T+1} + Q_{6j} + Q'_{6j} \}] Q'_{5j} \\
&+ \gamma_1 \sigma^3 \{ Q_2 \sum_{i,j=1}^p (I_{T+1} \circ (Q'_{4ij})) i Q'_{3ij} + Q_2 \sum_{i=1}^p (I_{T+1} \circ (Q'_{6i})) i Q'_{5i} \\
&\quad + \sum_{i,j=1}^p Q_{3ij} \{ (I_{T+1} \circ Q'_{4ij}) i \}' Q_2' + \sum_{i=1}^p Q_{5i} \{ (I_{T+1} \circ Q_{6i}) i \}' Q_2' \} \\
&+ \gamma_2 \sigma^4 \{ \sum_{i,j,l,m=1}^p Q_{3ij} Q'_{3lm} \text{tr}[Q'_{4lm} (I_{T+1} \circ Q_{4ij})] + \sum_{i,j,l=1}^p Q_{3ij} Q'_{5l} \text{tr}[Q'_{6l} (I_{T+1} \circ Q_{4ij})] \\
&\quad + \sum_{i,j,l=1}^p Q_{5i} Q'_{3jl} \text{tr}[Q'_{4jl} (I_{T+1} \circ Q_{6i})] + \sum_{i,j=1}^p Q_{5i} Q'_{5j} \text{tr}[Q'_{6j} (I_{T+1} \circ Q_{6i})] \}.
\end{aligned}$$

104 where " $\circ$ " is the Hadamard matrix product, and where  $Q_1 = \Gamma_2' A_1 + \Gamma_1' P_1$ ,  $Q_2 = \Gamma_2' A_2 + \Gamma_1' P_2$ ,  
105  $Q_{3ij} = \Gamma_2' A_{3ij}$ ,  $Q_{4ij} = A_{4ij}$ ,  $Q_{5i} = \Gamma_1' P_{3i}$  and  $Q_{6i} = P_{4i}$ .

106 *Proof.* See Appendix B.2.

107 The OLS bias is then given by

108 **Theorem 3.** *The bias in OLS regression of the ARX(p) model is*

$$E[\hat{\alpha}_i - \alpha_i] = \frac{1}{2} \{ \text{tr}(H_i |_{\delta} J) - \delta' H_i |_{\delta} \delta \} + o(T^{-1}),$$

109 where  $J$  is given in Lemma 2,  $\delta = Q_1 + \sigma^2 \sum_{i,j=1}^p Q_{3ij} \text{tr}(Q_{4ij}) + \sigma^2 \sum_{i=1}^p Q_{5i} \text{tr}(Q_{6i})$ , and where  
110 the Hessian is evaluated at  $\delta$ , or, equivalently, at  $E[Z'Z] = \bar{Z}' \bar{Z} + \sigma^2 \sum_{i,j=1}^p \Gamma_4' e_i e_j' \Gamma_4 \text{tr}(G_i' G_j)$  and at  
111  $E[\zeta] = P_1 + \sigma^2 \sum_{i=1}^p P_{3i} \text{tr}(P_{4i})$ .

112 *Proof.* See Appendix B.3.

113 The approximation in Theorem 3 is 'unfiltered', to use the terminology of [11], by which we mean  
114 there are likely to be unnecessary  $o(T^{-1})$  terms in the approximation that could in principle be removed.  
115 Doing this was found to be difficult, and would for example require a generalisation of the results in  
116 Lemma 2 below for traces of products of the matrices  $G_i$ ,  $i = 1, \dots, p$ . It may well be, as found in [15]

117 for the ARX(1), that no kurtosis coefficient  $\gamma_2$  need enter the approximation to order  $O(T^{-1})$ . In what  
 118 follows we specialise Theorem 3 to AR(1) cases with and without constant, remove the  $o(T^{-1})$  terms,  
 119 and agree with the results in [15] for these cases.

### 120 3.1. AR(1) models

121 We drop the subscript in  $G_1$  for notational convenience in the following: since  $p = 1$ , the matrices  
 122  $G_2, \dots, G_p$  are no longer used. Moreover, the form of the expected value  $J$  is the same as in Lemma  
 123 2, but we may remove the redundant summation over indices  $i, j, l$  and  $m$ . Some of the matrices used  
 124 above for the ARX( $p$ ) adjust straightforwardly when  $p = 1$ , for example

$$\Lambda = \begin{pmatrix} 1 & 0 & \cdot & \cdot & \cdot & 0 \\ -\lambda & 1 & & & & \cdot \\ 0 & -\lambda & 1 & & & \cdot \\ \cdot & \cdot & \cdot & \cdot & & \cdot \\ \cdot & \cdot & \cdot & \cdot & \cdot & \cdot \\ 0 & \cdot & \cdot & 0 & -\lambda & 1 \end{pmatrix} \quad \text{and} \quad \Omega = \begin{pmatrix} \omega & 0 & \cdot & \cdot & \cdot & 0 \\ 0 & 1 & 0 & & & \cdot \\ 0 & 0 & 1 & & & \cdot \\ & & & \cdot & & \cdot \\ \cdot & & & & \cdot & 0 \\ 0 & \cdot & \cdot & \cdot & 0 & 1 \end{pmatrix},$$

125 and the matrices  $B_i$  are as in Theorem 2 with  $N = k + 1$ . Some specialisations for others are collected  
 126 in Corollary 1 below.

127 **Corollary 1.** *The matrices  $P_1, \dots, P_4$  and  $A_1, \dots, A_4$  in Lemma 1 specialise to the following when  $p = 1$ :*

$$\begin{aligned} A_1 &= \text{vec}(\bar{Z}'\bar{Z}), \quad A_2 = (e_1 \otimes \bar{Z}'G) + (\bar{Z}'G \otimes e_1), \quad A_3 = e_1 \otimes e_1, \quad A_4 = G'G \\ P_1 &= \text{vec}[\bar{Z}'(\lambda\bar{y}_{-1} + X\beta)], \quad P_2 = \{(\lambda\bar{y}_{-1} + X\beta)'G\} \otimes e_1 + \bar{Z}'(\lambda G + \Gamma_3), \\ P_3 &= e_1, \quad P_4 = G'(\lambda G + \Gamma_3), \end{aligned}$$

#### 128 3.1.1. AR(1) with known mean

129 [8] and [7] found the bias in estimation of the model with  $p = 1$  and  $k = 0$  to be  $-\frac{2\lambda}{T}$  to order  
 130  $O(T^{-1})$  under a normality assumption. [15] show that the  $O(T^{-1})$  bias for this model is the same when  
 131 the disturbance terms are skewed with non-zero excess kurtosis. Here we confirm that skewness and  
 132 kurtosis in the error terms does not affect the bias to order  $O(T^{-1})$ . To do this we need to specialise  
 133 Theorem 3 to the pure AR(1) case, and we need to filter out any unnecessary  $o(T^{-1})$  terms, only keeping  
 134 the  $O(T^{-1})$  part of the bias.

135 As in [11], the matrix  $G$  can be written as  $G = [\omega F' : C]$ , where  $F' = (1, \lambda, \lambda^2, \dots, \lambda^{T-1})$  and

$$C = \begin{pmatrix} 0 & 0 & \cdot & \cdot & \cdot & 0 \\ 1 & 0 & & & & \cdot \\ \lambda & 1 & 0 & 0 & & \cdot \\ \cdot & \cdot & \cdot & \cdot & & \cdot \\ \cdot & \cdot & \cdot & \cdot & \cdot & \cdot \\ \lambda^{T-2} & \cdot & \cdot & \lambda & 1 & 0 \end{pmatrix}.$$



136 In order to obtain filtered results for the AR(1) with known mean, certain products of  $G$ ,  $G'$ ,  $C$ ,  
 137  $C'$  and  $F$  must be replaced with their  $O(T)$  or  $O(1)$  approximations, otherwise the resulting OLS  
 138 bias approximation is accurate to  $O(T^{-1})$  but includes some unnecessary  $o(T^{-1})$  terms. The required  
 139 approximations are summarised in

140 **Lemma 3.** (See [15] and [11])

$$\begin{aligned} \text{tr}(C'C) &= T\left(\frac{1}{1-\lambda^2}\right) + o(T), \quad \text{tr}(CC'C) = T\lambda\left(\frac{1}{1-\lambda^2}\right)^2 + o(T) \\ \text{tr}(C'CC'C) &= T(\lambda^2 + 1)\left(\frac{1}{1-\lambda^2}\right)^3 + o(T), \\ \text{tr}(G'G) &= T\left(\frac{1}{1-\lambda^2}\right) + \frac{\omega}{1-\lambda^2} - \left(\frac{1}{1-\lambda^2}\right)^2 + o(1) \\ \text{tr}(GG'C) &= T\lambda\left(\frac{1}{1-\lambda^2}\right)^2 + o(T), \quad \text{tr}(GG'GG') = \text{tr}(C'CC'C) + o(T) \end{aligned}$$

141 A further Lemma is also required:

142 **Lemma 4.** In the AR(1) model with known mean, i.e.  $y = \lambda y_{-1} + u$ , the matrices  $P_1, \dots, P_4$  and  $A_1, \dots, A_4$   
 143 specialise to the following:

$$\begin{aligned} A_1 &= 0, \quad A_2 = 0, \quad A_3 = 1, \quad A_4 = G'G \\ P_1 &= 0, \quad P_2 = 0, \quad P_3 = 1, \quad P_4 = G'(\lambda G + [0_{T \times 1} : I_T]), \end{aligned}$$

144 giving

$$Q_1 = 0, \quad Q_2 = 0, \quad Q_3 = e_1, \quad Q_4 = A_4, \quad Q_5 = e_2, \quad Q_6 = P_4.$$

145 Here  $e_1$  and  $e_2$  are  $2 \times 1$  unit matrices with unity in rows 1 and 2 respectively.

146 *Proof.* Follows from Corollary 1 by setting  $k = 0$ .

147 The following can now be stated:

148 **Corollary 2.** The bias in OLS estimation of the autoregressive coefficient in the pure AR(1) model, with  
 149  $\gamma_1$  and  $\gamma_2$  taking any finite value, is

$$E[\hat{\lambda} - \lambda] = -\frac{2\lambda}{T} + o(T^{-1}).$$

150 *Proof.* See Appendix B.4.

151 This agrees with the [15] generalisation of the original [8] and [7] result. As a by-product, the following  
 152 approximation for the evaluated Hessian matrix was found, where the subscript  $i$  is dropped given that  
 153 there is only one element in  $\delta$ :

$$H_i|_\delta = \begin{pmatrix} \frac{2\lambda(1-\lambda^2)^2}{T^2\sigma^4} & -\frac{(1-\lambda^2)^2}{T^2\sigma^4} \\ -\frac{(1-\lambda^2)^2}{T^2\sigma^4} & 0 \end{pmatrix} + o(T^{-2}).$$

## 154 3.1.2. AR(1) with unknown mean

155 A similar reduction can be done for the case where  $p = 1$  and the model includes a constant, see [8]  
 156 and [7] for the original approximation under normality. Unlike in the case with known mean, it cannot  
 157 be seen immediately that the four  $\gamma_1$  terms in Lemma 2 sum to zero. In order to find this, it is necessary  
 158 to use the following results

159 **Lemma 5.** (See [11] and [15], respectively)

$$i'_t F F' F = \left(\frac{1}{1-\lambda}\right)\left(\frac{1}{1-\lambda^2}\right) + o(T), \quad i'_T C (I_T \circ C' C) i_T = \frac{T}{(1-\lambda^2)(1-\lambda)} + o(T)$$

160 **Corollary 3.** *The bias in OLS estimation of the autoregressive coefficient in the AR(1)+constant model,*  
 161 *with  $\gamma_1$  and  $\gamma_2$  taking any finite value, is*

$$E[\hat{\lambda} - \lambda] = -\frac{1+3\lambda}{T} + o(T^{-1}).$$

162 *Proof.* See Appendix B.5.

163 Here we confirm that having skewness or kurtosis in the error terms does not affect the bias in these  
 164 cases to order  $O(T^{-1})$ , see [15]. Note that the results here assume a random covariance-stationary  
 165 startup, while [15] and [12] assume a fixed startup. [11] show that this distinction has an effect on the  
 166 bias in the case of normal errors to order  $O(T^{-2})$  but not to order  $O(T^{-1})$ . Here we can see that there  
 167 is still no effect to order  $O(T^{-1})$  when the errors are non-normal. It is also clear that the skewness  
 168 parameter  $\gamma_1$  does not enter the  $O(T^{-1})$  bias expression for the more general AR( $p$ ) with no intercept:  
 169 in this case  $\bar{Z} = \bar{y}_{-1} = 0$  so  $Q_2 = 0$  from Lemmas 1 and 2. An earlier sequence of papers starting with  
 170 [23] and ending with [24] also shows this, and shows the same for  $\gamma_2$ , though the validity conditions  
 171 for these approximations include finite error moments up to the 12th and 16th orders for the known and  
 172 unknown mean cases, respectively.

173 **4. Conclusion**

174 The results here facilitate the direct application of the validity framework outlined in [2]. The approach  
 175 is applicable in principle to models with both dynamics and endogeneity. The methodology has been  
 176 used to obtain an order  $O(T^{-1})$  bias expression for the OLS coefficient estimator of a general ARX( $p$ )  
 177 model with non-normal disturbances. Moreover, we have confirmed the non normality results in [15] for  
 178 the AR(1) with and without constant. It may be possible in the future to extend the methodology to other  
 179 estimators such as 2SLS.

180 **Conflicts of Interest**

181 The author declares no conflict of interest.

182 **References**

- 183 1. Phillips, G. An Alternative Approach to Obtaining Nagar-Type moment Approximations in  
184 Simultaneous Equation models. *Journal of Econometrics* **2000**, 97(2), 345–364.
- 185 2. Kiviet, J.; Phillips, G. Improved variance estimation of maximum likelihood estimators in  
186 stable first-order dynamic regression models. *Computational Statistics and Data Analysis* **2014**,  
187 76, 424–448.
- 188 3. Neudecker, H. Some Theorems on Matrix Differentiation with Special Reference to Kronecker  
189 Matrix Products. *Journal of the American Statistical Association* **1969**, 64, pp. 953–963.
- 190 4. Neudecker, H.; Wansbeek, T. Some Results on Commutation Matrices, with Statistical  
191 Applications. *The Canadian Journal of Statistics / La Revue Canadienne de Statistique* **1983**,  
192 11, pp. 221–231.
- 193 5. Magnus, J.; Neudecker, H. The commutation matrix: some properties and applications. *The*  
194 *Annals of Statistics* **1979**, 7(2), 318–394.
- 195 6. Magnus, J.; Neudecker, H. *Matrix differential calculus with applications in statistics and*  
196 *econometrics (Revised edition)*; Wiley, 2002.
- 197 7. Marriott, F.; Pope, J. Bias in the estimation of autocorrelations. *Biometrika* **1954**, 61, 393–403.
- 198 8. Kendall, M. Note on bias in the estimation of autocorrelation. *Biometrika* **1954**, 61, 403–404.
- 199 9. Nagar, A. The bias and moment matrix of the general k-class estimators of the parameters in  
200 simultaneous equations. *Econometrica* **1959**, 27, 575–595.
- 201 10. Kiviet, J.; Phillips, G. Moment approximation for least-squares estimators in dynamic regression  
202 models with a unit root. *Econometrics Journal* **2005**, 8, 1–28.
- 203 11. Kiviet, J.; Phillips, G. Higher-order asymptotic expansions of the least-squares estimation bias  
204 in first-order dynamic regression models. *Computational Statistics and Data Analysis* **2012**,  
205 56(11), 3706–3729.
- 206 12. Bao, Y. The approximate moments of the least squares estimator for the stationary autoregressive  
207 model under a general error distribution. *Econometric Theory* **2007**, 23(05), 1013–1021.
- 208 13. Kiviet, J.; Phillips, G. The bias of the ordinary least squares estimator in simultaneous equation  
209 models. *Economics Letters* **1996**, 53, 161–167.
- 210 14. Rilestone, P.; Srivastava, S.; Ullah, A. The second-order bias and mean squared error of nonlinear  
211 estimators. *Journal of Econometrics* **1996**, 75, 369–395.
- 212 15. Bao, Y.; Ullah, A. The second-order bias and mean-squared error of estimators in time-series  
213 models. *Journal of Econometrics* **2007**, 140(2), 650–669.
- 214 16. Sargan, J. The validity of Nagar's expansion for the moments of econometric estimators.  
215 *Econometrica* **1974**, 42, 169–176.
- 216 17. Sargan, J. Econometric estimators and the Edgeworth expansion. *Econometrica* **1976**,  
217 44, 421–448.
- 218 18. Phillips, P. Approximations to some finite sample distributions associated with a first-order  
219 stochastic difference equation. *Econometrica* **1977**, 45, 463–485.
- 220 19. Shao, J.; Tu, D. *The Jackknife and Bootstrap*, first ed.; Springer-Verlag: 175th Avenue, New  
221 York, 1995.

- 222 20. Shao, J. On resampling methods for variance and bias estimation in linear models. *Annals of*  
 223 *Statistics* **1988**, *16*, 986–1008.
- 224 21. Kiviet, J.; Phillips, G. Bias assessment and reduction in linear error-correction models. *Journal*  
 225 *of Econometrics* **1994**, *63*, 215–243.
- 226 22. Ullah, A. *Finite Sample Econometrics*; Oxford University Press, 2004.
- 227 23. Bhansali, R. Effects of not knowing the order of an autoregressive process on the mean squared  
 228 error of prediction-I. *Journal of the American Statistical Association* **1981**, *76*, 588–597.
- 229 24. Shaman, P.; Stine, R. The bias of autoregressive coefficient estimators. *Journal of The American*  
 230 *Statistical Association* **1988**, *83*, 842–848.

## 231 A. Theorem 2

### 232 A.1. Derivation of $H_i$

233  $H_i$  can be found using the Second Identification Theorem in [6]<sup>3</sup>. This requires the second differential  
 234 of  $\hat{\alpha}_i = e_i'(Z'Z)^{-1}\hat{\zeta}$  to be expressed in the form  $(d\hat{\delta})'A_i(d\hat{\delta})$  where  $A_i$  is a constant matrix. The Hessian  
 235 is then  $H_i = \frac{1}{2}(A + A')$ . The first differential can be calculated as follows:

$$\begin{aligned} d\hat{\alpha} &= (d(Z'Z)^{-1})\hat{\zeta} + (Z'Z)^{-1}d\hat{\zeta} \\ &= -(\hat{\zeta}' \otimes I_N)(Z'Z)^{\otimes(-2)}\text{vec}(d(Z'Z)) + (Z'Z)^{-1}d\hat{\zeta}. \end{aligned}$$

236 We can write  $\text{vec}(\hat{\zeta}) = \Gamma_1\hat{\delta}$  and  $\text{vec}(Z'Z) = \Gamma_2\hat{\delta}$ . Using this gives

$$d\hat{\alpha} = Nd\hat{\delta},$$

237 where  $N = (Z'Z)^{-1}\Gamma_1 - (\hat{\zeta}' \otimes I_N)(Z'Z)^{\otimes(-2)}\Gamma_2$ . For the first differential we now have

$$\begin{aligned} d\hat{\alpha} &= \text{vec}(I_N N(d\hat{\delta})) \\ &= ((d\hat{\delta})' \otimes I_N)\text{vec}N, \end{aligned}$$

238 which is a convenient form for calculating the second:

$$d^2\hat{\alpha}_i = d(\text{vec}(N))'((d\hat{\delta}) \otimes I_N)e_i.$$

239 In the above we use  $d(d\hat{\delta}') = 0$ , since  $d\hat{\delta}$  is the constant vector increment in the differential  $d\hat{\alpha}$ . Note  
 240 that the term  $((d\hat{\delta}) \otimes I_N)e_i$  can be written as  $B_i\hat{\delta}$ , where  $B_i$  is derived in part (ii) below, so that

---

<sup>3</sup> In particular, see the second line of Table 1 in Chapter 10.

$$d^2 \hat{\alpha}_i = (d\text{vec}(N))' B_i (d\hat{\delta}).$$

241 The remaining task is to put the second differential in the form

$$d^2 \hat{\alpha}_i = (d\hat{\delta})' M B_i (d\hat{\delta})$$

242 for some  $M$ , then the Hessian can be identified as  $H_i = \frac{1}{2} (M B_i + B_i' M')$ . We therefore need to put  
243  $d\text{vec}(N)$  in the form  $d\text{vec}(N) = M' d\hat{\delta}$ .

244 From  $N = (Z'Z)^{-1} \Gamma_1 - (\hat{\zeta}' \otimes I_N)(Z'Z)^{\otimes(-2)} \Gamma_2$  we have

$$d\text{vec}(N) = d\text{vec}[(Z'Z)^{-1} \Gamma_1] - d\text{vec}[(\hat{\zeta}' \otimes I_N)(Z'Z)^{\otimes(-2)} \Gamma_2].$$

245 The first term of this can be written as  $-(\Gamma_1' \otimes I_N)(Z'Z)^{\otimes(-2)} \Gamma_2 d\hat{\delta}$ , and the second term can be written  
246 as follows:

$$\begin{aligned} & \text{vec}((\hat{\zeta}' \otimes I_N)(Z'Z)^{\otimes(-2)} \Gamma_2 + (\hat{\zeta}' \otimes I_N) d[(Z'Z)^{\otimes(-2)} \Gamma_2]) \\ &= (((Z'Z)^{\otimes(-2)} \Gamma_2)' \otimes I_N)(I_N \otimes V_1) \text{vec}(d\hat{\zeta}') \\ &\quad - [\Gamma_2' \otimes (\hat{\zeta}' \otimes I_N)] \text{vec}((Z'Z)^{\otimes(-2)} d((Z'Z)^{\otimes 2})(Z'Z)^{\otimes(-2)}) \\ &= (((Z'Z)^{\otimes(-2)} \Gamma_2)' \otimes I_N)(I_N \otimes V_1) \Gamma_1 \hat{\delta} \\ &\quad - [\Gamma_2' \otimes (\hat{\zeta}' \otimes I_N)] (Z'Z)^{\otimes(-4)} [(I_N \otimes V_2) + (V_3 \otimes I_N)] \Gamma_2 d\hat{\delta}, \end{aligned}$$

247 where, following the result in the exercise on p48 of [6],  $V_1 = (K_{21} \otimes I_2)(I_1 \otimes \text{vec}(I_N)) = (K_{N1} \otimes$   
248  $I_N) \text{vec}(I_N) = \text{vec}(I_N)$ ,  $V_2 = (K_{NN} \otimes I_N)(I_N \otimes \text{vec}(Z'Z))$  and  $V_3 = (I_N \otimes K_{NN})(\text{vec}(Z'Z) \otimes I_N)$ .

249 From the above we have

$$d\text{vec}(N) = M' d\hat{\delta},$$

250 therefore  $d^2 \hat{\alpha}_i = (d\hat{\delta})' M B_i (d\hat{\delta})$  and the Hessian is

$$H_i = \frac{1}{2} (M B_i + B_i' M').$$

251 **A.2. Derivation of  $B_i$**   $\{(\hat{\delta} \otimes I_N)\}$  is an  $N^2(N+1) \times N$  matrix and  $e_i$  is an  $N \times 1$  vector with unity

252 in element  $i$ . Let  $d\hat{\delta} = (d\hat{r}_1, \dots, d\hat{r}_{N(N+1)})'$ . Then we have the  $N^2(N+1) \times 1$  vector  $\{(\hat{\delta} \otimes I_N)\} e_i =$   
253  $(d\hat{r}_1 e_i, \dots, d\hat{r}_{N(N+1)} e_i) = B_i d\hat{\delta}$  for some constant  $N^2(N+1) \times N(N+1)$  matrix  $B_i$ . Consider the case  
254  $i = 1$ , and let the  $nm$ -th element of a matrix  $A$  be denoted as  $(A)_{nm}$ . Here we see that  $(B_1)_{nm} = 1$  for  
255  $n = 1$  and  $m = 1$ , and for  $n = N+1$  and  $m = 2$ , and more generally for  $n = qN+1$  and  $m = q+1$   
256 up to  $q = N^2 + N - 1$ , with all other elements being zero. Similarly, we can see that  $(B_2)_{nm} = 1$  for  
257  $n = qN+2$  and  $m = q+1$ , up to  $q = N^2 + N - 1$ , and zero otherwise. For general  $i$  we therefore have  
258  $(B_i)_{nm} = 1$  for  $n = qN+i$ ,  $m = q+1$  and  $q = 1, \dots, N^2 + N - 1$ .

259 **B. ARX models**260 *B.1. Lemma 1*

261 We have

$$\begin{aligned}
\text{vec}(\hat{\zeta}) &= \text{vec}(Z'y) \\
&= \text{vec}\left[\left(\bar{Z} + \sum_{i=1}^p G_i v e_i \Gamma_4\right)' \left\{ \left( \sum_{j=1}^p \lambda_j \bar{y}_{-j} \right) + X\beta + \left\{ \left( \sum_{j=1}^p \lambda_j G_j \right) + \Gamma_3 \right\} v \right\} \right] \\
&= \text{vec}\left[ \bar{Z}' \left\{ \left( \sum_{j=1}^p \lambda_j \bar{y}_{-j} \right) + X\beta \right\} \right] + \left\{ \bar{Z}' \left\{ \left( \sum_{j=1}^p \lambda_j G_j \right) + \Gamma_3 \right\} \right. \\
&\quad \left. + \sum_{i=1}^p \left[ \left\{ \left( \sum_{j=1}^p \lambda_j \bar{y}_{-j} \right) + X\beta \right\}' G_i \right] \otimes \Gamma_4' e_i \right\} v + \sum_{i=1}^p \Gamma_4' e_i v' G_i' \left\{ \left( \sum_{j=1}^p \lambda_j G_j \right) + \Gamma_3 \right\} v \right] \\
&= P_1 + P_2 v + P_3 v' P_4 v.
\end{aligned}$$

262 Similarly,

$$\begin{aligned}
\text{vec}(Z'Z) &= \text{vec}\left[\left(\bar{Z} + \sum_{i=1}^p G_i v e_i \Gamma_4\right) \left(\bar{Z} + \sum_{i=1}^p G_i v e_i \Gamma_4\right)'\right] \\
&= \text{vec}(Z'Z) + \sum_{i=1}^p \left\{ \left( \Gamma_4' e_i \otimes \bar{Z}' G_i \right) + \left( \bar{Z}' G_i \otimes \Gamma_4' e_i \right) \right\} v + \sum_{i,j=1}^p \left( \Gamma_4' e_j \otimes \Gamma_4' e_i \right) v' G_i' G_j v \\
&= A_1 + A_2 v + \sum_{i,j=1}^p A_{3ij} v' A_{4ij} v.
\end{aligned}$$

263 *B.2. Lemma 2*

264 Since  $\hat{\delta} = \Gamma_2' \text{vec}(Z'Z) + \Gamma_1' \text{vec}(\hat{\zeta})$  we have  $\hat{\delta} = Q_1 + Q_2 v + \sum_{i,j=1}^p Q_{3ij} v' Q_{4ij} v + Q_{5i} v' Q_{6i} v$  from  
265 Lemma 1 where  $Q_1 = \Gamma_2' A_1 + \Gamma_1' P_1$ ,  $Q_2 = \Gamma_2' A_2 + \Gamma_1' P_2$ ,  $Q_{3ij} = \Gamma_2' A_{3ij}$ ,  $Q_{4ij} = A_{4ij}$  and  $Q_{5i} = \Gamma_1' P_{3i}$ ,  
266  $Q_{6i} = P_{4i}$ . To calculate  $J$  we can write

$$\begin{aligned}
\hat{\delta} \hat{\delta}' &= \left( Q_1 + Q_2 v + \sum_{i,j=1}^p Q_{3ij} v' Q_{4ij} v + \sum_{i=1}^p Q_{5i} v' Q_{6i} v \right) \\
&\quad \times \left( Q_1 + Q_2 v + \sum_{i,j=1}^p Q_{3ij} v' Q_{4ij} v + \sum_{i=1}^p Q_{5i} v' Q_{6i} v \right)',
\end{aligned}$$

267 then after eliminating terms that have zero expectation we have

$$\begin{aligned}
E[\hat{\delta}\hat{\delta}'] &= Q_1Q_1' + Q_1 \sum_{i,j=1}^p E[v'Q_4'v]Q_{3ij}' + Q_1 \sum_{i=1}^p E[v'Q_6'v]Q_{5i}' + Q_2E[vv']Q_2' \\
&+ Q_2 \sum_{i,j=1}^p E[vv'Q_{4ij}'v]Q_{3ij}' + Q_2 \sum_{i=1}^p E[vv'Q_{6i}'v]Q_{5i}' + \sum_{i,j=1}^p Q_{3ij}E[v'Q_{4ij}v]Q_1' \\
&+ \sum_{i,j=1}^p Q_{3ij}E[v'Q_{4ij}vv']Q_2' + \sum_{i,j,k,l=1}^p Q_{3ij}E[v'Q_{4ij}vv'Q_{4lm}'v]Q_{3lm}' + \sum_{i,j,l=1}^p Q_{3ij}E[v'Q_{4ij}vv'Q_{6l}'v]Q_{5l}' \\
&+ \sum_{i=1}^p Q_{5i}E[v'Q_{6i}v]Q_1' + \sum_{i=1}^p Q_{5i}E[v'Q_{6i}vv']Q_2' + \sum_{i,j,l=1}^p Q_{5i}E[v'Q_{6i}vv'Q_{4jl}'v]Q_{3jl}' \\
&+ \sum_{i,j=1}^p Q_{5i}E[v'Q_{6i}vv'Q_{6j}'v]Q_{5j}'.
\end{aligned}$$

268 In order to calculate these expected values, the results in Appendix A.5 of [22] are used. Doing this  
269 gives

$$\begin{aligned}
J &= Q_1Q_1' + \sigma^2 \sum_{i,j=1}^p \text{tr}(Q_{4ij}')Q_1Q_{3ij}' + \sigma^2 \sum_{i=1}^p \text{tr}(Q_{6i}')Q_1Q_{5i}' + \sigma^2Q_2Q_2' \\
&+ \sigma^3\gamma_1Q_2 \sum_{i,j=1}^p (I_{T+1} \circ (Q_{4ij}'))iQ_{3ij}' + \sigma^3\gamma_1Q_2 \sum_{i=1}^p (I_{T+1} \circ (Q_{6i}'))iQ_{5i}' \\
&+ \sigma^2 \sum_{i,j=1}^p \text{tr}(Q_{4ij})Q_{3ij}Q_1' + \sigma^3\gamma_1 \sum_{i,j=1}^p Q_{3ij}\{(I_{T+1} \circ Q_{4ij}')i\}'Q_2' \\
&+ \sigma^4 \sum_{i,j,k,l=1}^p Q_{3ij}\text{tr}[Q_{4lm}'\{\gamma_2(I_{T+1} \circ Q_{4ij}) + \text{tr}(Q_{4ij})I_{T+1} + Q_{4ij} + Q_{4ij}'\}]Q_{3lm}' \\
&+ \sigma^4 \sum_{i,j,l=1}^p Q_{3ij}\text{tr}[Q_{6l}'\{\gamma_2(I_{T+1} \circ Q_{4ij}) + \text{tr}(Q_{4ij})I_{T+1} + Q_{4ij} + Q_{4ij}'\}]Q_{5l}' \\
&+ \sigma^2 \sum_{i=1}^p \text{tr}(Q_{6i})Q_{5i}Q_1' + \sigma^3\gamma_1 \sum_{i=1}^p Q_{5i}\{(I_{T+1} \circ Q_{6i}')i\}'Q_2' \\
&+ \sigma^4 \sum_{i,j,l=1}^p Q_{5i}\text{tr}[Q_{4jl}'\{\gamma_2(I_{T+1} \circ Q_{6i}) + \text{tr}(Q_{6i})I_{T+1} + Q_{6i} + Q_{6i}'\}]Q_{3jl}' \\
&+ \sigma^4 \sum_{i,j=1}^p Q_{5i}\text{tr}[Q_{6j}'\{\gamma_2(I_{T+1} \circ Q_{6i}) + \text{tr}(Q_{6i})I_{T+1} + Q_{6i} + Q_{6i}'\}]Q_{5j}'.
\end{aligned}$$

### 270 B.3. Theorem 3

271 From Theorem 1 we have  $E[\hat{\alpha}_i - \alpha_i] = \frac{1}{2}(\text{tr}(H_i|\delta J) - \delta'H_i|\delta\delta) + o(T^{-1})$ . From Lemma 2 we  
272 have  $\delta = E[Q_1 + Q_2v + Q_3v'Q_4v + Q_5v'Q_6v] = Q_1 + Q_3\text{tr}(Q_4) + Q_5\text{tr}(Q_6)$ . The Hessian  $H_i$   
273 was found in Theorem 2, and to evaluate it note that  $E[Z'Z] = E[(\bar{Z} + G_bv\Gamma_4)'(\bar{Z} + G_bv\Gamma_4)]$  and  
274  $E[\zeta] = E[(\bar{Z} + G_bv\Gamma_4)'(\sum_{i=1}^p \lambda_i\bar{y}_{-i}) + X\beta + [(\sum_{i=1}^p \lambda_iG_i)\Gamma_3]v]$ . These two expected values are

275 then calculated in the same way as Lemma 2.

276

#### 277 B.4. Corollary 2

278 (i) *Specialising the matrices  $H_i$  and  $H_i|_\delta$*

279 Recall that  $H_i = MB_i + B'_iM'$ , or  $H = 2BM'$  for the case at hand since  $H$ ,  $M$  and  $B$  are all scalars.

280 We also have  $B = 1$ , so that  $H = M'$ . In the matrix representation of  $M'$  we have the following for the

281 AR(1) with known mean:

$$M' = [e_1 \otimes (\hat{\zeta}' \otimes I_1)](Z'Z)^{\otimes(-4)}[(I_1 \otimes V_2) + (V_3 \otimes I_1)]e'_1 \\ - (e_2 \otimes I_1)(Z'Z)^{\otimes(-2)}e'_1 - [(Z'Z)^{\otimes(-2)}e'_1]' \otimes I_1](I_1 \otimes V_1)e'_2.$$

282 It follows that

$$H = e_1\hat{\zeta}'(Z'Z)^{\otimes(-4)}(V_2 + V_3)e'_1 - e_2(Z'Z)^{\otimes(-2)}e'_1 - e_1(Z'Z)^{\otimes(-2)}V_1e'_2.$$

283 Noting the specialisations  $V_1 = I_1$  and  $V_2 = V_3 = \text{vec}(Z'Z)$ , the expression for  $H$  reduces to

$$H = 2e_1\hat{\zeta}'(Z'Z)^{\otimes(-4)}\text{vec}(Z'Z)e'_1 - e_2(Z'Z)^{\otimes(-2)}e'_1 - e_1(Z'Z)^{\otimes(-2)}e'_2.$$

284 Finally, recall that  $Z'Z$  and  $\hat{\zeta}$  are scalars, so that

$$H = \begin{pmatrix} 2\hat{\zeta}(Z'Z)^{-3} & -(Z'Z)^{-2} \\ -(Z'Z)^{-2} & 0 \end{pmatrix} \text{ and } H|_\delta = \begin{pmatrix} 2\zeta(E[Z'Z])^{-3} & -(E[Z'Z])^{-2} \\ -(E[Z'Z])^{-2} & 0 \end{pmatrix}.$$

285 The expected values in  $H|_\delta$  are  $E[Z'Z] = \sigma^2 \text{tr}(G'G)$  and  $\hat{\zeta} = \sigma^2 \lambda \text{tr}(G'G)$  from Theorem

286 2, and from Lemma 3 the largest terms in each are  $O(T)$ . More specifically, we have

287  $\text{tr}(G'G) = (\frac{T}{1-\lambda^2}) + (\frac{\omega}{1-\lambda^2}) - (\frac{1}{1-\lambda^2})^2 + \frac{\lambda^{2T}}{1-\lambda^2}(\frac{1}{1-\lambda^2} - \omega)$ . This means that the non-zero elements

288 in  $H|_\delta$  are at most  $O(T^{-2})$ , though there will also be some smaller  $o(T^{-2})$  contributions due to

289 the  $o(T)$  components of  $E[Z'Z]$  and  $\hat{\zeta}$ . In both  $H|_\delta$  and  $J$  we can discard contributions of order

290  $O(\lambda^{sT})$  for  $s > 0$  since there no (explosive)  $O(\lambda^{-sT})$  contributions in either. Any products in

291  $H|_\delta J$  involving  $O(\lambda^{sT})$  terms will be  $o(T^{-1})$ . Therefore we use the approximation of  $\text{tr}(G'G)$  up

292 to order  $O(1)$  that appears in Lemma 3, omitting the term  $\frac{\lambda^{2T}}{1-\lambda^2}(\frac{1}{1-\lambda^2} - \omega)$ . After simplifying  $H|_\delta$  we have

293

$$H|_\delta = \tilde{H}|_\delta + o(T^{-2}), \\ \text{where } \tilde{H}|_\delta = \begin{pmatrix} \frac{2\lambda(1-\lambda^2)^2}{\sigma^4 T^2} & -\frac{(1-\lambda^2)^2}{\sigma^4 T^2} \\ -\frac{(1-\lambda^2)^2}{\sigma^4 T^2} & 0 \end{pmatrix}.$$



294 (i) Specialising the matrix  $J$

295 When  $p = 1$  we have

$$\begin{aligned}
J &= Q_1 Q_1' + \sigma^2 \text{tr}(Q_4') Q_1 Q_3' + \sigma^2 \text{tr}(Q_6') Q_1 Q_5' + \sigma^2 Q_2 Q_2' \\
&+ \sigma^3 \gamma_1 Q_2 (I_{T+1} \circ (Q_4')) i Q_3' + \sigma^3 \gamma_1 Q_2 (I_{T+1} \circ (Q_6')) i Q_5' \\
&+ \sigma^2 \text{tr}(Q_4) Q_3 Q_1' + \sigma^3 \gamma_1 Q_3 \{(I_{T+1} \circ Q_4') i\}' Q_2' \\
&+ \sigma^4 Q_3 \text{tr}[Q_4' \{\gamma_2 (I_{T+1} \circ Q_4) + \text{tr}(Q_4) I_{T+1} + Q_4 + Q_4'\}] Q_3' \\
&+ \sigma^4 Q_3 \text{tr}[Q_6' \{\gamma_2 (I_{T+1} \circ Q_4) + \text{tr}(Q_4) I_{T+1} + Q_4 + Q_4'\}] Q_5' \\
&+ \sigma^2 \text{tr}(Q_6) Q_5 Q_1' + \sigma^3 \gamma_1 Q_5 \{(I_{T+1} \circ Q_6) i\}' Q_2' \\
&+ \sigma^4 Q_5 \text{tr}[Q_4' \{\gamma_2 (I_{T+1} \circ Q_6) + \text{tr}(Q_6) I_{T+1} + Q_6 + Q_6'\}] Q_3' \\
&+ \sigma^4 Q_5 \text{tr}[Q_6' \{\gamma_2 (I_{T+1} \circ Q_6) + \text{tr}(Q_6) I_{T+1} + Q_6 + Q_6'\}] Q_5'
\end{aligned}$$

296 where the redundant summation over  $i, j, l$  and  $m$  has been removed. In Lemma 4 we see that  $Q_2 = 0$   
297 in the pure AR(1) case, therefore all the terms in  $\gamma_1$  here are zero. The terms in  $\gamma_2$  are not all zero, and  
298 to consider their combined influence on the bias we can use the decomposition  $J = J_1 + J_2$  where the  
299 terms in  $\gamma_2$  are collected in  $J_2$ :

$$\begin{aligned}
J_1 &= Q_1 Q_1' + \sigma^2 \text{tr}(Q_4') Q_1 Q_3' + \sigma^2 \text{tr}(Q_6') Q_1 Q_5' \\
&+ \sigma^2 \text{tr}(Q_4) Q_3 Q_1' + \sigma^4 Q_3 \text{tr}[Q_4' \{\text{tr}(Q_4) I_{T+1} + Q_4 + Q_4'\}] Q_3' \\
&+ \sigma^4 Q_3 \text{tr}[Q_6' \{\text{tr}(Q_4) I_{T+1} + Q_4 + Q_4'\}] Q_5' \\
&+ \sigma^2 \text{tr}(Q_6) Q_5 Q_1' + \sigma^4 Q_5 \text{tr}[Q_4' \{\text{tr}(Q_6) I_{T+1} + Q_6 + Q_6'\}] Q_3' \\
&+ \sigma^4 Q_5 \text{tr}[Q_6' \{\text{tr}(Q_6) I_{T+1} + Q_6 + Q_6'\}] Q_5' \\
J_2 &= \sigma^4 \gamma_2 \{ Q_3 Q_3' \text{tr}[Q_4' (I_{T+1} \circ Q_4)] + Q_3 Q_5' \text{tr}[Q_6' (I_{T+1} \circ Q_4)] \\
&+ Q_5 Q_3' \text{tr}[Q_4' (I_{T+1} \circ Q_6)] + Q_5 Q_5' \text{tr}[Q_6' (I_{T+1} \circ Q_6)] \}.
\end{aligned}$$

300 We have  $Q_2 = 0$ ,  $Q_1 = 0$ ,  $Q_3 = e_1$ ,  $Q_4 = A_4$ ,  $Q_5 = e_2$  and  $Q_6 = P_4$ , which enables further  
301 specialisation:

$$\begin{aligned}
J_1 &= \sigma^4 [\{\text{tr}(A_4)\}^2 + 2\text{tr}(A_4' A_4)] e_1 e_1' + \sigma^4 \{\text{tr}(A_4) \text{tr}(P_4') + 2\text{tr}(P_4' A_4)\} e_1 e_2' \\
&+ \sigma^4 \{\text{tr}(P_4) \text{tr}(A_4) + 2\text{tr}(A_4' P_4)\} e_2 e_1' + \sigma^4 [\{\text{tr}(P_4)\}^2 + \text{tr}(P_4' P_4)] \\
&+ \text{tr}(P_4' P_4)] e_2 e_2' \\
J_2 &= \sigma^4 \gamma_2 \{\text{tr}[A_4' (I_{T+1} \circ A_4)] e_1 e_1' + \text{tr}[P_4' (I_{T+1} \circ A_4)] e_1 e_2' \\
&+ \text{tr}[A_4' (I_{T+1} \circ P_4)] e_2 e_1' + \text{tr}[P_4' (I_{T+1} \circ P_4)] e_2 e_2'\}.
\end{aligned}$$

302 The next task is to make this more explicit in terms of  $\lambda$ . Moreover, since we wish to discard all  
303  $o(T^{-1})$  terms from the product  $H|_\delta J$ , and since the largest terms in  $H|_\delta$  are  $O(T^{-2})$ , we must discard all  
304  $o(T)$  terms from  $J$ . To do this, recall from Lemma 4 that  $A_4 = G'G$  and  $P_4 = G'(\lambda G + [0_{T \times 1} : I_T])$  and

305 that we have approximations for the traces of products of these in Lemma 3. Let  $\tilde{J}$ ,  $\tilde{J}_1$  and  $\tilde{J}_2$  denote,  
 306 respectively, the versions of  $J$ ,  $J_1$ ,  $J_2$  where  $o(T)$  terms are excluded.

307 We have the following for the first term in  $\tilde{J}_1$ :

$$\begin{aligned} \{tr(A_4)\}^2 + 2tr(A_4'A_4) &= \{tr(G'G)\}^2 + 2tr(G'GG'G) \\ &= \left\{T\left(\frac{1}{1-\lambda^2}\right)\right\}^2 + 2\left\{\frac{\omega}{1-\lambda^2} - \left(\frac{1}{1-\lambda^2}\right)^2\right\}\left(\frac{T}{1-\lambda^2}\right) \\ &\quad + 2T(\lambda^2 + 1)\left(\frac{1}{1-\lambda^2}\right)^3 + o(T) \\ &= \frac{T(2 + T - (T-2)\lambda^2)}{(1-\lambda^2)^3} + o(T) \end{aligned}$$

308 We have the following for the second and third terms in  $\tilde{J}_1$ :

$$\begin{aligned} tr(A_4)tr(P_4') + 2tr(P_4'A_4) &= tr(G'G)tr((\lambda G + [0_{T \times 1} : I_T])'G) \\ &\quad + 2tr((\lambda G + [0_{T \times 1} : I_T])'GG'G) \\ &\quad + tr\{(\omega F : C)(\omega F : C)'(\omega F : C)[0_{T \times 1} : I_T]'\} \\ &= \lambda\{tr(G'G)\}^2 + 2[\lambda tr(G'GG'G) \\ &\quad + tr\{(\omega F : C) \begin{pmatrix} \omega^2 F'F & \omega F'C \\ \omega C'F & C'C \end{pmatrix} \begin{pmatrix} 0_{1 \times T} \\ I_T \end{pmatrix}\}] \\ &= \lambda\left(\frac{T}{1-\lambda^2}\right)^2 + 2\lambda\left\{\frac{\omega}{1-\lambda^2} - \left(\frac{1}{1-\lambda^2}\right)^2\right\}\left(\frac{T}{1-\lambda^2}\right) \\ &\quad + 2T(\lambda^2 + 1)\left(\frac{1}{1-\lambda^2}\right)^3 + T\lambda\left(\frac{1}{1-\lambda^2}\right)^2\} + o(T) \\ &= \frac{T\lambda(4 + T(1-\lambda^2))}{(1-\lambda^2)^3} + o(T) \end{aligned}$$

309 For the fourth term in  $\tilde{J}_1$  we have

$$\{tr(P_4)\}^2 + tr(P_4'P_4) + tr(P_4'P_4') = \lambda^2\{tr(G'G)\}^2 + tr(P_4'P_4) + tr(P_4'P_4').$$

310 We consider  $tr(P_4'P_4)$  and  $tr(P_4'P_4')$  individually now:

$$\begin{aligned} tr(P_4'P_4) &= tr\{(\lambda G + [0_{T \times 1} : I_T])'GG'(\lambda G + [0_{T \times 1} : I_T])\} \\ &= \lambda^2 tr(G'GG'G) + \lambda tr([0_{T \times 1} : I_T]'GG'G) + \lambda tr([0_{T \times 1} : I_T]'GG'G) \\ &\quad + tr\{[0_{T \times 1} : I_T]'(\omega F : C)(\omega F : C)'[0_{T \times 1} : I_T]\} \\ &= \lambda^2 T(\lambda^2 + 1)\left(\frac{1}{1-\lambda^2}\right)^3 + 2\lambda(T\lambda\left(\frac{1}{1-\lambda^2}\right)^2) + T\left(\frac{1}{1-\lambda^2}\right) + o(T) \end{aligned}$$

$$\begin{aligned}
tr(P_4'P_4) &= tr\{(\lambda G + [0_{T \times 1} : I_T])'G(\lambda G + [0_{T \times 1} : I_T])'G\} \\
&= \lambda^2 tr(G'GG'G) + 2\lambda tr\{[0_{T \times 1} : I_T]'GG'G\} \\
&\quad + tr([0_{T \times 1} : I_T]'(\omega F : C)[0_{T \times 1} : I_T]'(\omega F : C)) \\
&= \lambda^2 T(\lambda^2 + 1)\left(\frac{1}{1 - \lambda^2}\right)^3 + 2\lambda\{T\lambda\left(\frac{1}{1 - \lambda^2}\right)^2\} + o(T).
\end{aligned}$$

311 This gives

$$\{tr(P_4)\}^2 + tr(P_4'P_4) + tr(P_4'P_4) = \frac{T\{(T + 4)\lambda^2 - (T + 1)\lambda^4 + 1\}}{(1 - \lambda^2)^3},$$

312 so that the final form of  $\tilde{J}_1$  is as follows:

$$\begin{aligned}
J_1 &= \sigma^4 \left\{ \frac{T(2 + T - (T - 2)\lambda^2)}{(1 - \lambda^2)^3} \right\} e_1 e_1' \\
&\quad + \sigma^4 \left\{ \frac{T\lambda(4 + T(1 - \lambda^2))}{(1 - \lambda^2)^3} \right\} e_1 e_2' \\
&\quad + \sigma^4 \left\{ \frac{T\lambda(4 + T(1 - \lambda^2))}{(1 - \lambda^2)^3} \right\} e_2 e_1' \\
&\quad + \sigma^4 \left\{ \frac{T\{(T + 4)\lambda^2 - (T + 1)\lambda^4 + 1\}}{(1 - \lambda^2)^3} \right\} e_2 e_2' + o(T).
\end{aligned}$$

313 We can do a similar specialisation for  $J_2$ , using  $tr[Q_6(I_{T+1} \circ Q_4)] = \lambda \sum_{i=1}^{T+1} (G'G)_{ii}^2$  and  $tr[Q_4'(I_{T+1} \circ$   
314  $Q_4)] = \sum_{i=1}^{T+1} (G'G)_{ii}^2$ . We have

$$J_2 = \gamma_2 \sigma^4 \begin{pmatrix} \sum_{i=1}^{T+1} (G'G)_{ii}^2 & \lambda \sum_{i=1}^{T+1} (G'G)_{ii}^2 \\ \lambda \sum_{i=1}^{T+1} (G'G)_{ii}^2 & tr[P_4'(I_{T+1} \circ P_4)] \end{pmatrix}.$$

315 Here it is unnecessary to find  $\tilde{J}_2$ , because it can already be shown (below) that  $tr(\tilde{H}|_\delta J_2) = 0$  to  
316 order  $O(T^{-1})$ .

317

318 (iii) *Finding the bias result*

319 The relevant elements of  $\tilde{H}|_\delta J_2$  are

320

$$\begin{aligned}
(\tilde{H}_i|_\delta J_2)_{11} &= \frac{2\lambda(1 - \lambda^2) \sum_{i=1}^{T+1} (G'G)_{ii}}{T^2 \sigma^4} - \frac{\lambda(1 - \lambda^2) \sum_{i=1}^{T+1} (G'G)_{ii}}{T^2 \sigma^4} \\
&\quad + o(T^{-1})
\end{aligned}$$

$$\text{and } (\tilde{H}_i|_\delta J_2)_{22} = -\frac{\lambda(1 - \lambda^2) \sum_{i=1}^{T+1} (G'G)_{ii}}{T^2 \sigma^4} + o(T^{-1}),$$

321 showing that  $tr(\tilde{H}|_\delta J_2)$  is  $o(T^{-1})$ .

322 To complete the proof, we see that  $tr(\tilde{H}|_\delta \tilde{J}_1)$  simplifies to  $\frac{-2\lambda}{T}$  and, using  $\delta = \sigma^2 tr(G'G)e_1 +$   
 323  $\sigma^2 \lambda tr(G'G)e_2$  from Theorem 3, it is straightforward to show that  $\delta' \tilde{H}|_\delta \delta$  is zero.

324

### 325 B.5. Corollary 3

326 (i) *Specialising the matrices  $H_i$  and  $H_i|_\delta$*

327 For the Corollary we are only interested in the  $i = 1$  case. We have  $B_1 =$   
 328  $(e_1, 0, e_2, 0, e_3, 0, e_4, 0, e_5, 0, e_6, 0)'$  where the  $e_i$  here are  $12 \times 1$  vectors with 1 in position  $i$  and 0  
 329 elsewhere. The matrix  $M$  is written in terms of  $\Gamma_1, \Gamma_2, \hat{\zeta}, E[Z'Z], V_2$  and  $V_3$ . These can be specialised

330 as follows:  $E[Z'Z] = \bar{Z}'\bar{Z} + \Gamma_4'\Gamma_4 tr(G'G) = \begin{pmatrix} T\left(\frac{\beta}{1-\lambda}\right)^2 + tr(G'G) & T\left(\frac{\beta}{1-\lambda}\right) \\ T\left(\frac{\beta}{1-\lambda}\right) & T \end{pmatrix}$

$$V_2 = (K_{22} \otimes I_2)(I_2 \otimes vec(E[Z'Z])) = \begin{pmatrix} T\left(\frac{\beta}{1-\lambda}\right)^2 + tr(G'G) & 0 \\ T\left(\frac{\beta}{1-\lambda}\right) & 0 \\ 0 & T\left(\frac{\beta}{1-\lambda}\right)^2 + tr(G'G) \\ 0 & T\left(\frac{\beta}{1-\lambda}\right) \\ T\left(\frac{\beta}{1-\lambda}\right) & 0 \\ T & 0 \\ 0 & T\left(\frac{\beta}{1-\lambda}\right) \\ 0 & T \end{pmatrix}$$

332  $V_3 = V_2$

333  $\hat{\zeta} = P_1 + P_3 tr(P_4) = \lambda \begin{pmatrix} T\left(\frac{\beta}{1-\lambda}\right)^2 \\ \frac{T\beta}{1-\lambda} \end{pmatrix} + \beta \begin{pmatrix} \frac{T\beta}{1-\lambda} \\ T \end{pmatrix} + \beta \begin{pmatrix} 1 \\ 0 \end{pmatrix} (\lambda tr(G'G) + tr(G'\Gamma_3)).$

334

335 (i) *Specialising the matrix  $J$*

336 The matrix  $J$  is in terms of  $Q_1$  to  $Q_6$ . These are

337  $Q_1 = \begin{pmatrix} T(\beta/(1-\lambda))^2 \\ T\beta/(1-\lambda) \\ T\beta/(1-\lambda) \\ T \\ T\{\lambda(\beta/(1-\lambda))^2 + \beta^2/(1-\lambda)\} \\ T\{\lambda\beta/(1-\lambda) + \beta\} \end{pmatrix}, Q_2 = \begin{pmatrix} (2\beta/(1-\lambda))i'G \\ i'G \\ i'G \\ ((2\lambda\beta/(1-\lambda) + \beta))i'G + (\beta/(1-\lambda))i'\Gamma_3 \\ \lambda i'G + i'\Gamma_3 \end{pmatrix}$

338  $Q_3 = (1, 0, 0, 0, 0, 0)'$ ,  $Q_4 = G'G$ ,  $Q_5 = (0, 0, 0, 0, 1, 0)'$ ,

339  $Q_6 = \lambda G'G + G'[0_{T \times 1} : I_T]$ .

340

341 These were found using

$$342 A_1 = \begin{pmatrix} T(\beta/(1-\lambda))^2 \\ T\beta/(1-\lambda) \\ T\beta/(1-\lambda)T \end{pmatrix}, A_2 = \begin{pmatrix} 2\bar{y}'_{-1}G \\ i'G \\ i'G \\ 0 \end{pmatrix}, A_3 = \begin{pmatrix} 1 \\ 0 \\ 0 \\ 0 \end{pmatrix}, A_4 = G'G,$$

$$343 P_1 = \lambda \begin{pmatrix} T(\beta/(1-\lambda))^2 \\ T\beta/(1-\lambda) \end{pmatrix} + \beta \begin{pmatrix} T\beta/(1-\lambda) \\ T \end{pmatrix}$$

$$344 P_2 = \begin{pmatrix} (\lambda\bar{y}_{-1} + i\beta)'G \\ 0 \end{pmatrix} + \lambda Z'G + \bar{Z}'\Gamma_3, P_3 = (1, 0)', P_4 = \lambda G'G + G'\Gamma_3.$$

345  
346 (iii) *Finding the bias result*

347 To find the result in Corollary 3 it is necessary to evaluate  $H|_\delta$  using  $\Gamma_1, \Gamma_2, \hat{\zeta}, E[Z'Z], V_2$  and  $V_3$ , and  
348 to calculate  $J$  for the specific versions of  $Q_1$  to  $Q_6$  in (ii). Due to the term  $E[Z'Z]^{\otimes(-4)}$  in particular,  
349 it seemed necessary to use a symbolic programming tool: a Mathematica programme is available from  
350 the author's website, where the final reduction is made. A program for Corollary 2 is also available. As  
351 before, we must use the  $O(T)$  approximations in Lemma 3 in order to filter out any smaller  $o(T^{-1})$  terms  
352 in the approximation. To deal with the terms in  $\gamma_1$  we have

$$353 Q_2(I_{T+1} \circ (Q'_4))iQ'_3 = A(1, 0, 0, 0, 0, 0), Q_2(I_{T+1} \circ (Q'_6))iQ'_5 = \lambda A(0, 0, 0, 0, 1, 0),$$

$$354 Q_3\{(I_{T+1} \circ Q'_4)i\}'Q'_2 = \{A(1, 0, 0, 0, 0, 0)\}', Q_5\{(I_{T+1} \circ Q_6)i\}'Q'_2 = \lambda\{A(0, 0, 0, 0, 1, 0)\}' \text{ where}$$

$$355 A = \begin{pmatrix} (2\beta/(1-\lambda))A_1 \\ A_1 \\ A_1 \\ 0 \\ ((2\lambda\beta/(1-\lambda)) + \beta)A_1 + (\beta/(1-\lambda))tr(C'C) \\ \lambda A_1 + tr(C'C) \end{pmatrix}$$

356 and  $A_1 = i'_T\omega^3 F F' F + i'_T C (I_T \circ C' C) i_T$ . We then use the results in Lemma 4 as before, along with  
357 those in Lemma 5.

358 © April 19, 2015 by the author; submitted to *Econometrics* for possible open access  
359 publication under the terms and conditions of the Creative Commons Attribution license  
360 <http://creativecommons.org/licenses/by/4.0/>.